Biostatistics I: Linear Regression

Multiple Linear Regression

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The Multiple Linear Regression Model

Basic assumptions:

- **single** continuous **response** variable
- **multiple covariates** of mixed type (continuous or categorical)

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The model is then formally written as:

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y_i = \beta_0 + \beta_1 x_{i1} + \beta_2 x_{i2} + \ldots + \beta_p x_{ip} + \underbrace{\varepsilon_i}_{\text{additive linear systematic component}}\\ \text{ (linear predictor)}\\ \text{ E}(\varepsilon_i) = 0, \quad \text{var}(\varepsilon_i) = \sigma^2, \quad i = 1, \ldots, n
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$$
\left(\begin{array}{cc} \bullet & \bullet & \bullet \end{array}\right)
$$

- Extension of simple linear regression to multiple covariates.
- **Note:** Both are **univariate** models!

What Makes the Linear Model Linear?

A **linear** regression model is **linear in the regression coefficients** and the error term.

- $y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \epsilon$
- $\mathbf{y} = \beta_0 + \beta_1 \mathbf{x}_1^2 + \beta_2 \log(\mathbf{x}_2) + \boldsymbol{\varepsilon}$
- $\log(y) = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \varepsilon$

Linear Not linear

• $y = \beta_0 + \exp(\beta_1 x_1 + \beta_2 x_2) + \varepsilon$ $\bullet \textbf{ y } = \beta_0 + \beta_1 \textbf{x}_1 / (\beta_2 \textbf{x}_2) + \boldsymbol{\varepsilon}$

•
$$
\mathbf{y} = \beta_0 + \beta_1 \mathbf{x}_1^{\beta_2} + \boldsymbol{\varepsilon}
$$

Example: Child Growth

Our data might look like this:

How would our regression model look like,

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How would our regression model look like,

$$
\textrm{height}_i = \beta_0 + \beta_1 \textrm{age}_i + \beta_2 \textrm{sex}_i + \beta_3 \textrm{race}_i + \varepsilon_i?
$$

Coefficients of Continuous Covariates

In the model

$$
\textrm{height}_i = \beta_0 + \beta_1 \textrm{age}_i + \beta_2 \textrm{sex}_i + \beta_3 \textrm{race}_i + \varepsilon_i
$$

 β_1 describes the change in the expected height when \texttt{age} is $\texttt{increased}$ by \texttt{one} **unit** and all **other covariates are held constant**.

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$$
\begin{aligned} \mathrm{height}_{age} &= \beta_0 + \ \beta_1 \mathrm{age} + \beta_2 \mathrm{sex} + \beta_3 \mathrm{race} \\ \mathrm{height}_{age+1} &= \beta_0 + \ \beta_1 (\mathrm{age} + 1) \ + \beta_2 \mathrm{sex} + \beta_3 \mathrm{race} \end{aligned}
$$

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\begin{array}{l} \mathrm{height}_{age} = \beta_0 + \ \beta_1 \mathrm{age} + \beta_2 \mathrm{sex} + \beta_3 \mathrm{race} \\ \\ \mathrm{height}_{age+1} = \beta_0 + \ \beta_1 (\mathrm{age} + 1) + \beta_2 \mathrm{sex} + \beta_3 \mathrm{race} \\ \\ \mathrm{height}_{age+1} - \mathrm{height}_{age} = \beta_1 (\mathrm{age} + 1) - \beta_1 \mathrm{age} & = \ \beta_1 \end{array}
$$

This results in the linear predictors:

$$
\begin{aligned}\n\text{boy (sex=0):} \quad & \beta_0 + \beta_1 \text{age} + \beta_3 \text{race} \\
\text{girl (sex=1):} \quad & \beta_0 + \beta_1 \text{age} + \beta_2 + \beta_3 \text{race}\n\end{aligned}
$$

What would this look like for the **effect of race**?

caucasian (race = 0): $\beta_0 + \beta_1$ **age** + β_2 **sex** asian (race = 1): $\quad \beta_0 + \beta_1 \mathrm{age} + \beta_2 \mathrm{sex} + \beta_3$ other (race = 2): $\beta_0 + \beta_1$ age + β_2 sex + $2\beta_3$

To **avoid the link** between effects of different categories we need **additional parameters**.

In general: One parameter less than the number of categories.

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Most common coding choices:

Dummy coding

Effect coding

Dummy Coding

Using **dummy coding**, the model is:

$$
\text{height}_i = \beta_0 + \beta_1 \text{age}_i + \beta_2 \text{sex}_i + \beta_3 \text{race}_i^{(asian)} + \beta_4 \text{race}_i^{(other)} + \varepsilon_i
$$

This leads to the following linear predictors:

caucasian:
$$
\beta_0 + \beta_1 \text{age} + \beta_2 \text{sex} + \beta_3 0 + \beta_4 0 = \beta_0 + \beta_1 \text{age} + \beta_2 \text{sex}
$$

\nasian: $\beta_0 + \beta_1 \text{age} + \beta_2 \text{sex} + \beta_3 1 + \beta_4 0 = \beta_0 + \beta_1 \text{age} + \beta_2 \text{sex} + \beta_3$
\nother: $\beta_0 + \beta_1 \text{age} + \beta_2 \text{sex} + \beta_3 0 + \beta_4 1 = \beta_0 + \beta_1 \text{age} + \beta_2 \text{sex} + \beta_4$

Dummy Coding

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Effect Coding

Using **effect coding**, the model is:

$$
\text{height}_i = \beta_0 + \beta_1 \text{age}_i + \beta_2 \text{sex}_i + \left\vert \beta_3 \text{race}_i^{(1)} \right\vert + \left\vert \beta_4 \text{race}_i^{(2)} \right\vert + \varepsilon_i
$$

Effect coding will lead to the following linear predictors:

caucasian:
$$
\beta_0 + \beta_1 \text{age} + \beta_2 \text{sex} + \beta_3 1 + \beta_4 0 = \beta_0 + \beta_1 \text{age} + \beta_2 \text{sex} + \beta_3
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\nasian: $\beta_0 + \beta_1 \text{age} + \beta_2 \text{sex} + \beta_3 0 + \beta_4 1 = \beta_0 + \beta_1 \text{age} + \beta_2 \text{sex} + \beta_4$
\nother: $\beta_0 + \beta_1 \text{age} + \beta_2 \text{sex} + \beta_3 (-1) + \beta_4 (-1) = \beta_0 + \beta_1 \text{age} + \beta_2 \text{sex} - \beta_3 - \beta_4$

Effect Coding

Dummy coding:

$$
\beta_0 + \beta_1 \text{age} + \beta_2 \text{sex} + \beta_3 \text{race}^{(asian)}_i + \beta_4 \text{race}^{(other)}_i
$$

In dummy coding, the intercept β_0 is the expected outcome when **all covariate values are zero**, i.e., for a caucasian (race^(asian) = race^(other) = 0) boy (sex = 0) with zero years of age.

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Effect coding:

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\beta_0 + \beta_1 \text{age} + \beta_2 \text{sex} + \beta_3 \text{race}^{(1)} + \beta_4 \text{race}^{(2)}
$$

With effect coding there is no scenario where all effects are zero.

In **effect coding** the intercept represents the **average expected response over all categories** (when all other covariates are zero).

$$
\begin{aligned} \text{height}_{cauc.} &= \beta_0 + \beta_3\\ \text{height}_{asian} &= \beta_0 + \beta_4\\ \text{height}_{other} &= \beta_0 - \beta_3 - \beta_4 \end{aligned}
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$$

$$
\frac{\text{height}_{cauc.} + \text{height}_{asian} + \text{height}_{other}}{3} = \frac{\beta_0 + \beta_3 + \beta_0 + \beta_4 + \beta_0 - \beta_3 - \beta_4}{3} \\ = \frac{3\beta_0}{3} = \beta_0
$$

Multiple Linear Regression in Matrix Notation

The basic model of **multiple linear regression in matrix notation** is

$$
\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\varepsilon}, \quad \mathrm{E}(\boldsymbol{\varepsilon}) = \mathbf{0}, \quad \mathrm{var}(\boldsymbol{\varepsilon}) = \sigma^2 \mathbf{I} \qquad \qquad \mathrm{I} =
$$

$$
= \begin{pmatrix} 1 & 0 & \dots & 0 \\ 0 & 1 & \dots & 0 \\ \vdots & & \ddots & \vdots \\ 0 & 0 & \dots & 1 \end{pmatrix}
$$

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$$

$$
\begin{pmatrix}y_1\\ \vdots\\ y_n\end{pmatrix}=\begin{pmatrix}1&x_{11}&\dots&x_{1p}\\ \vdots&\vdots&&\vdots\\ 1&x_{n1}&\dots&x_{np}\end{pmatrix}\begin{pmatrix}\beta_0\\ \vdots\\ \beta_p\end{pmatrix}+\begin{pmatrix}\varepsilon_1\\ \vdots\\ \varepsilon_n\end{pmatrix}
$$

Sidenote: Tansposing Vectors and Matrices

$$
\mathbf{y} = \begin{pmatrix} y_1 \\ \vdots \\ y_n \end{pmatrix} \quad \Rightarrow \quad \mathbf{y}^\top = (y_1, \ldots, y_n)
$$

Sidenote: Tansposing Vectors and Matrices

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\mathbf{y} = \begin{pmatrix} y_1 \\ \vdots \\ y_n \end{pmatrix} \quad \Rightarrow \quad \mathbf{y}^\top = (y_1, \ldots, y_n)
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$$
\mathbf{X} = \begin{pmatrix} 1 & x_{11} & \ldots & x_{1p} \\ \vdots & \vdots & & \vdots \\ 1 & x_{n1} & \ldots & x_{np} \end{pmatrix} \quad \Rightarrow \quad \mathbf{X}^{\top} = \begin{pmatrix} 1 & \ldots & 1 \\ x_{11} & \ldots & x_{n1} \\ \vdots & & \vdots \\ x_{1p} & \ldots & x_{np} \end{pmatrix}
$$

Estimation via OLS

Ordinary Least Squares (OLS) Estimator

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Ordinary Least Squares (OLS) Estimator

$$
\sum_{i=1}^n (y_i - \mathbf{x}_i^\top \boldsymbol\beta)^2 \longrightarrow \min_{\hat{\varepsilon}}\limits_{}
$$

The **least squares principle** in matrix notation

$$
(\mathbf{y} - \mathbf{X}\boldsymbol{\beta})^{\top}(\mathbf{y} - \mathbf{X}\boldsymbol{\beta}) \longrightarrow \min_{\boldsymbol{\beta}}
$$