# **Biostatistics I: Linear Regression**

#### Hypothesis Tests & Model Fit

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#### **Linear Regression Model:**

$$y_i = \mathbf{x}_i^ op oldsymbol{eta} + arepsilon_i, \quad \mathrm{E}(arepsilon_i) = 0, \quad \mathrm{var}(arepsilon_i) = \sigma^2$$

#### **Estimation via OLS:**

$$\hat{oldsymbol{eta}} = (\mathbf{X}^{ op}\mathbf{X})^{-1}\mathbf{X}^{ op}\mathbf{y} \qquad ext{and} \qquad \hat{\sigma}^2 = rac{1}{n-p-1} oldsymbol{\hat{arepsilon}}^{ op} oldsymbol{\hat{arepsilon}}$$

$$\operatorname{height}_i = \beta_0 + \beta_1 \operatorname{age}_i + \beta_2 \operatorname{sex}_i + \varepsilon_i$$



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How confident are we that the observed difference is "real"?



Because we have a (random) sample of the data, there is **always some difference**, even when the true effect is zero.

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likely • When we know the distribution of  $\hat{oldsymbol{eta}}$  we can calculate how (un)likely the observed difference is if there is no effect. unlikely no effect

#### **Distribution of the OLS estimates:**

$$\mathsf{lf}\, \varepsilon_i \sim N(0,\sigma^2) \texttt{:} \qquad \qquad \boldsymbol{\hat{\beta}} \sim N\!\left(\boldsymbol{\beta}, \underbrace{\sigma^2(\mathbf{X}^\top \mathbf{X})^{-1}}_{\mathrm{var}(\boldsymbol{\hat{\beta}})}\right)$$

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#### Standardized effect with estimated variance $\hat{\sigma}_j$ :

$$rac{\hat{eta}_j - eta_j}{\hat{\sigma}_j} \sim t(n-p-1), \hspace{1em} j = 0, \dots, p$$

 $eta_j$  is the (assumed) true value,  $\sigma_j$  is the standard deviation of  $\hat{eta}_j$ 

t(n-p-1) is the **Student's** *t*-distribution with n-p-1 degrees of freedom.

#### **Research Question:**

Does  $\mathbf{x}_j$  contribute (significantly) to the model (i.e., explain variation in  $\mathbf{y}$ )?

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#### **Corresponding hypothesis:**

$$H_0:eta_j=0, \quad H_1:eta_j
eq 0.$$

In general:

$$H_0:eta_j=eta_{0j},\quad H_1:eta_j
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The test statistic is the standardized regression coefficient

$$T_j = rac{{\hateta}_j - {eta}_{0j}}{{\hat\sigma}_j}, \quad j=0,\ldots,p.$$

null hypothesis $H_0:eta_j=eta_{0j}\ H_0:eta_j=eta_{0j}\ H_0:eta_j=eta_{0j}$ 

alternative hypothesis

 $egin{aligned} H_1:eta_j
eqeta_{0j}\ H_1:eta_j<eta_{0j}\ H_1:eta_j>eta_{0j} \end{aligned}$ 

rejection if $|T_j| > t_{1-lpha/2}(n-p-1) \ T_j < -t_{1-lpha}(n-p-1) \ T_j > t_{1-lpha}(n-p-1)$ 



- n = 108
- p = 2
- df = 108 2 1 = 105

• lpha=0.05

$$\bullet \ -t_{1-\alpha/2}=-1.98$$

• 
$$t_{1-lpha/2}=$$
 1.98

•  $\hat{eta}_2 = -2.14$ •  $\hat{\sigma}_2 = 0.94$ •  $\beta_{02} = 0$ •  $T_2 = \frac{-2.14}{0.94} = -2.27$ 



⇒ Reject the null hypothesis that sex has no effect on height.

### **Confidence Interval**

The (two-sided) (1-lpha)100% confidence interval for  $eta_{0j}$  can be calculated as

$$\left[ {\hat eta}_j - {\hat \sigma}_j t_{1-lpha/2} (n-p-1), \quad {\hat eta}_j + {\hat \sigma}_j t_{1-lpha/2} (n-p-1) 
ight]$$

by solving 
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by solving 
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 for  $eta_{0j}.$ 

#### In our example:

The 95% confidence interval for  $\beta_2$  is

[-2.14 - 0.94 imes 1.98, -2.14 + 0.94 imes 1.98] = [-4.01, -0.27]

#### **P-value**

The **p-value** is the probability to obtain the observed parameter estimate or a more extreme value (into the direction of  $H_1$ ) **under the null-hypothesis**.



$$p=2\min\{\Pr(t\leq T\mid H_0), \; \Pr(t\geq T\mid H_0)\}$$

In the example: p=2 imes 0.0127=0.0253

### **Example: Child Growth (smaller sample)**

How do things change if we had a **smaller sample**?

•  $n = 32 \Rightarrow df = 29$ •  $t_{1-\alpha/2} = 2.05$ •  $\hat{\beta}_2 = -3.21$ •  $\hat{\sigma}_2 = 1.97$ 

•  $T_2 = \frac{-3.21}{1.97} = -1.63$ 



⇒ Do not reject the null hypothesis that sex has no effect on height.

### **Interpretation of Test Results**

Example with n = 108:

"There is a difference in height between boys and girls."

Example with n = 32:

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#### **Possible phrasing:**

- "... height was associated with sex ..." or "... girls were 2.14cm shorter than boys ..."
- "... there was no evidence for an association between height and sex ..."
- "... we did not find an association between height and sex ..."

#### Do NOT use:

- "... there was no association/effect/difference ..."
- "... we found a non-significant association ..."
- "... with a trend towards significance ..."

### **Model Fit**

How much of the variation in  $\mathbf{y}$  is explained by the model?



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How much of the variation in  ${f y}$  is explained by the model?



total variation

residual variation

#### explained variation



covariate



covariate





#### **Overall-F-Test (Goodness of Fit Test)**

Simultaneous test for all regression coefficients:

$$H_0: eta_1=eta_2=\ldots=eta_p=0, \qquad H_1: eta_j
eq 0 ext{ for at least one } j.$$

The **test statistic** of the Goodness of fit test is:

$$F = \frac{\text{ESS}}{\text{RSS}} \frac{n - p - 1}{p}$$

Under  $H_0$ :

$$F\sim F(p,n-p-1)$$

 $\Rightarrow$  Reject the null hypothesis if  $F>F_{1-lpha}(p,n-p-1).$ 

### **Overall-F-Test (Goodness of Fit Test)**

The test statistic F and sums of squares are often shown in an **analysis of variance** table:

	variation	degrees of freedom	mean squared error	test statistic
explained variation	ESS	p	$ ext{MSE} = rac{ ext{ESS}}{p}$	$F = rac{ ext{MSE}}{ ext{MSR}}$
residual variation	RSS	n-p-1	$\mathrm{MSR} = rac{\mathrm{RSS}}{n-p-1}$	
total variation	TSS	n-1		

$$R^2 = rac{\sum (\hat{y}_i - ar{y})^2}{\sum (y_i - ar{y})^2} = rac{\mathrm{ESS}}{\mathrm{TSS}} = 1 - rac{\mathrm{RSS}}{\mathrm{TSS}}$$

Hence:  $0 \leq R^2 \leq 1$ 

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Special case for  $y = \beta_0 + \beta_1 x + \varepsilon$ :

 $R^2 = r_{xy}^2$  (Pearson correlation coefficient)

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Hence:  $0 \leq R^2 \leq 1$ 

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In multiple linear regression:

$$R^2 = r_{y\hat{y}}^2$$

 ${\it R}^2$  can only be used if

- models have the **same response** variable *y*,
- the number of regression coefficients is the same, and
- all models include an intercept.

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#### **Adjusted Coefficient of Determination**

To **correct for the size** of the model:

$$R_{adj}^2 = 1 - rac{n-1}{n-p-1}(1-R^2)$$